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2 Cold and Heat Related Mortality: A Cautionary Note on Current
3 Damage Functions with Net Benefits from Climate Change

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10

11 *Abstract*

12 Several economic assessments of climate change build on the assumption that reductions of cold-
13 related mortality will overcompensate increases in heat-related mortality at least for moderate levels
14 of global warming. Due to the lack of suitable epidemiological studies with sufficient spatial
15 coverage, many of these assessments rely on one particular dataset: projections of temperature-related
16 mortality in 17 countries published almost 20 years ago. Here, we reanalyse this dataset with a focus
17 on cardiovascular mortality, and present evidence for two flaws in the original analysis, which would
18 imply a significant bias towards finding net mortality benefits from climate change: i) the
19 combination of mortality data for all ages with data specific to the elderly, and ii) the confounding of
20 seasonal effects with direct temperature effects on mortality. This bias appears to be further amplified
21 in the integrated assessment models FUND and ENVISAGE, and related economic assessment tools
22 relying on the same calibration scheme, because heat-related cardiovascular mortality is assumed to
23 affect urban populations only in these models. In an exemplary calculation, we show that while
24 FUND currently projects a net reduction of approximately 380,000 deaths from cardiovascular
25 diseases globally per year at 1°C of global warming, correcting for the two potential flaws and
26 assuming equal vulnerability of urban and rural populations would result in a net increase of
27 cardiovascular mortality, with approximately 150,000 net additional deaths globally per year. Our
28 findings point to the urgent need of renewing damage functions on temperature-related mortality
29 currently applied in some of the most widely used integrated assessment models.

30 *Keywords:* climate change impacts, human health, temperature-related mortality, damage function,
31 integrated assessment model

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39 Recent economic analyses have shown that human health impacts may contribute disproportionately to
40 overall damage costs of climate change (e.g., Houser et al. 2015). Therefore, the specific form and
41 parameterization of damage functions to describe climate-sensitive health outcomes in these
42 assessments may critically influence the magnitude of estimated total damage costs.

43 Here, we are concerned with the epidemiological databasis on temperature-related mortality of one of
44 the most widely used integrated assessment models (IAMs) FUND. When Tol (2002a, b) first
45 integrated temperature-related mortality into FUND, it constituted a major step forward in integrated
46 assessment modelling. While many other IAMs work with relatively few, broad impact categories
47 (e.g., market and non-market impacts), FUND explicitly incorporates different sectors (i.e., human
48 health, agriculture, sea level rise, etc.), for which quantitative knowledge on climate impacts exist.

49 One of the challenges with such 'bottom-up' approaches has always been to find datasets at suitable
50 temporal and spatial resolution to construct damage functions, in particular for global-scale IAMs
51 such as FUND. Specifically, most of the epidemiological studies, projecting temperature-related
52 mortality under climate change, present their results only at the spatial resolution of small number of
53 individual cities, regions, or countries (see review by Huang et al. 2011). The functions used in FUND
54 to describe climate-induced changes in temperature-related mortality derive from a meta-analysis
55 published almost twenty years ago (Martens 1998). At the time this meta-analysis constituted one of
56 the few empirically based studies providing mortality projections with broader geographical scope
57 (i.e., 17 countries worldwide).

58 Here, we reanalyse the data presented by Martens (1998), with the objective to test the robustness of
59 one of Martens' major conclusions, namely that "[in most cities] global climate change is likely to
60 lead to a reduction in mortality rates due to decreasing winter mortality". Our reanalysis reveals a
61 number of questionable assumptions underlying Martens' results, including two potential flaws in the
62 data handling. We subsequently investigate how the bias in Martens' mortality projections that result
63 from these flaws propagates into FUND's empirical databases on temperature-related mortality. We
64 find that the assumptions made by Tol (2002a) in the extrapolation of Martens' data to FUND regions
65 further amplify the identified bias, contributing to the likely overestimating of net health benefits from
66 climate change in the model.

67 Our findings are relevant today, although they concern epidemiological data published almost two
68 decades ago, and they are relevant beyond FUND itself. The reason is that FUND has been used in
69 several recent economic assessments of climate damages, including latest assessments of the social
70 costs of carbon (SCC) for regulatory purposes in the UK and US (Fig. 1). Moreover, ENVISAGE
71 (Roson and van der Mensbrugge 2012), another broadly used IAM, ultimately relies on the same
72 data as FUND through a prominent economic assessment of climate impacts on human health
73 (Bosello et al. 2006) (Fig. 1). The latter together with original results of Martens (1998) have been
74 influential in shaping the public opinion about climate impact on heat and cold related mortality
75 (Lomborg 2007). Last but not least, due to the persistent lack of more up-to-date global estimates of
76 temperature-related mortality impacts (Huang et al. 2011), Tol's equations have even be used in recent
77 non-economic quantitative assessments of climate change impacts (Hayashi et al. 2010).

78 Our study goes beyond previous critique of the same data by Ackerman and Stanton (2008) by
79 providing additional quantitative evidence for the likely bias in estimates of temperature-related
80 mortality impacts underlying FUND, ENVISAGE and related economic assessments of climate

81 damage costs. Ackerman and Stanton (2008) relate this bias to three conceptual shortcomings in the
82 use of Martens' (1998) data by Tol (2002a) and Bosello et al. (2006): i) the lack of accounting for
83 acclimatization in temperature-mortality relationships, ii) missing out on extreme events by neglecting
84 daily temperature variability, and iii) the unsupported assumption that heat-related (cardiovascular)
85 mortality does not affect rural populations. We revisit their third point of criticism, but embed this in a
86 more detailed quantitative reanalysis of Martens (1998) and the use of his data in Tol (2002a).
87 Because of the flaws we believe to reveal in their original analyses, our study provides important
88 evidence that the update of damage formulations on temperature-related mortality in FUND and
89 related economic assessment tools is indispensable now.

90 *2 Material and Methods*

91 *2.1 Data Sources*

92 We retrieved the relevant source data on published temperature-mortality relationships from Table 2
93 of Martens (1998) (Online Resource 1 Table S1). Out of the full dataset presented by Martens (1998),
94 Tol (2002a) uses results for cardiovascular deaths in the age groups < and > 65 years, and heat-related
95 respiratory deaths in all ages. Our reanalysis only concerns cardiovascular mortality, because
96 projections for respiratory mortality are based on an extremely low sample size (one data point) and
97 because Martens (1998) only documents projections for cardiovascular mortality in detail. We took
98 the given projections of annual additional cardiovascular mortality (per 100,000 people) for a scenario
99 of approximately 1.2 °C increase in global mean temperature (GMT) in 17 countries together with the
100 corresponding age group specific mortality baselines from Table 3 of Martens (1998).

101 For the extrapolation of mortality projections to countries worldwide, we followed Tol (2002a) and
102 extracted the minimum and maximum mean monthly temperatures in capital cities from the (updated)
103 Leemans and Cramer (1991) database, which provides mean monthly temperatures on a global grid
104 based on 1931-1960 climatology (Online Resource 1 Table S5). To derive total additional annual
105 deaths in FUND regions, we used age-specific population data per country, for urban and total area,
106 taken from the UN Population Division (2014). Tol (2002a) does not specify the exact source of the
107 population data used in his analysis. As default we applied population data for the year 1990.

108 *2.2 Outline of reanalysis*

109 Our reanalysis consisted of the following four steps (Online Resource 1 Fig. S1).

110 *1. Reassessment of V-shaped relationship between ambient temperature and changes in*
111 *mortality.* We followed Martens (1998) assuming a V-shaped relationship between ambient
112 temperature and changes in mortality (Fig. 2), i.e. mortality rates increase linearly as temperatures
113 drop below or rise above a site-specific minimum mortality temperature (MMT). We calculated
114 combined effect estimates $\bar{\beta}$ (percent changes in mortality for a 1°C temperature change) as described
115 by Martens (1998), weighting individual data points β_j by their inverse variance (square of the
116 standard error s_j) as follows:

$$117 \quad \bar{\beta} = \frac{\sum \frac{1}{s_j^2} \beta_j}{\sum \frac{1}{s_j^2}} \quad (1)$$

118 where j indexes the epidemiological studies that entered Martens' meta-analysis (see Online Resource
 119 1 Table S1). Estimates were derived separately for the cold and warm range (below and above MMT),
 120 and age groups $<$ and $>$ 65 years: $\bar{\beta}_{cold,<65}$, $\bar{\beta}_{warm,<65}$, $\bar{\beta}_{cold,>65}$, $\bar{\beta}_{warm,>65}$.

121 These calculations uncovered two questionable assumptions in the meta-analysis of Martens (1998).
 122 First, Martens integrated all-ages data into age-specific estimates in order to increase sample sizes
 123 (Martens, personal communication). Yet, one would expect that this choice creates a significant bias
 124 in the age group $>$ 65 years, because older people are known to be especially vulnerable to
 125 temperature excursions from the comfort range, showing higher changes in mortality rates than the
 126 average population (Vardoulakis et al. 2014, Lee and Kim 2016). Second, the data points Martens
 127 extracted from Green et al. (1994) were suspiciously higher than any other data considered in the cold
 128 range (Fig. 2). Referring back to Green et al. (1994) showed that Martens considered simple
 129 differences between winter and summer mortality reported in this study (Online Resource 1 Table
 130 S2). Therefore, the data points from Green et al. (1994) entering Martens' meta-analysis obviously
 131 include seasonal effects not directly related to temperature, biasing results towards higher cold-related
 132 mortality (see e.g., Kinney et al. 2015 on the risk of confounding seasonal and temperature effects in
 133 the studies of temperature-related mortality).

134 To investigate the influence of these assumptions on Martens' mortality projections, we corrected for
 135 them one by one: i) excluding all ages data, ii) modifying estimates based on Green et al. (1994) by
 136 considering reported differences between cold and mild winters (aiming to isolate direct temperature
 137 effects; Online Resource 1 Table S3), and iii) applying both modifications together. These
 138 calculations resulted in a new set of revised estimates (denoted $\tilde{\beta}$ below) for each modification
 139 applied (Online Resource 1 Table S1).

140 2. *Revised projections of relative mortality changes for 17 countries.* Martens (1998) originally
 141 used monthly mean temperature projections for capital cities together with assumptions on city-
 142 specific MMTs (Online Resource 1 Table S4) to estimate changes in cardiovascular mortality due to
 143 climate change. Since we did not have access to these temperature projections we used the given data
 144 on country- and age-specific annual baseline mortalities ($B_{i,a}$), projected mortality changes ($C_{i,r,a}$)
 145 and weighted effect estimates ($\bar{\beta}_{r,a}$) to back-calculate an average annual measure of the applied local
 146 temperature changes (incorporating differences with respects to the city-specific MMTs and between
 147 the baseline and future climate), according to

$$148 \quad \overline{\Delta T}_{i,r} = \frac{C_{i,r,a}}{B_{i,a}\bar{\beta}_{r,a}} \quad (2)$$

149 where i is the country index, r indexes the temperature range (warm or cold), and a defines the age
 150 group ($<$ or $>$ 65 years) (the same indices are used in Eqs. 3-6). We chose to use data for age group $>$
 151 65 years only here, because of the small absolute mortality numbers in age group $<$ 65 years, which
 152 introduce large rounding errors (Online Resource 1 Fig. S2). Based on $\overline{\Delta T}_{i,r}$ and after rearranging Eq.
 153 2, revised annual mortality estimates $\tilde{C}_{i,r,a}$ can be calculated according to

$$154 \quad \tilde{C}_{i,r,a} = B_{i,a}\tilde{\beta}_{r,a}\overline{\Delta T}_{i,r} \quad (3)$$

155 We are aware that Eqs. 2 and 3 represent an extreme simplification compared to the standard method
 156 of deriving annual attributable mortality from temperature-mortality relationships at daily scale (e.g.,
 157 Vardoulakis et al. 2014). Therefore, we tested the validity of this approach for the purpose of our

158 study by reproducing part of the original mortality projections of Martens (1998) (Online Resource 1
159 Fig. S3).

160 3. *Extrapolation of mortality change estimates to countries worldwide.* We derived linear
161 regression equations linking country-specific, cold- and heat-related mortality projections in the two
162 age groups ($M_{i,r,a}$) with minimum and maximum mean monthly temperatures in capital cities ($T_{min,i}$,
163 $T_{max,i}$) as done by Tol (2002a):

$$164 \quad M_{i,cold,a} = p_0 + p_1 T_{min,i} \quad (4)$$

$$165 \quad M_{i,warm,a} = p_0 + p_1 T_{max,i} \quad (5)$$

166 where p_0 and p_1 are regression parameters. We calculated four different sets of regression equations
167 using i) original mortality projections of Martens (1998) ($C_{i,r,a}$), and ii) our revised projections
168 ($\tilde{C}_{i,r,a}$), correcting separately as well as simultaneously for the two potential flaws in Martens (1998).
169 Following Tol (2002a), we linearly rescaled given mortality projections to a 1°C increase in GMT
170 (relative to the 1.16°C GMT increase considered by Martens (1998)), such that, e.g., $M_{i,r,a} = f C_{i,r,a}$
171 with $f \approx 0.862$. These equations, in addition to those given in Tol (2002a) (see Online Resource 1
172 Table S6) were used to extrapolate mortality projections to countries worldwide, resulting in a dataset
173 comprising 188 countries.

174 4. *Calculation of absolute mortality estimates and aggregation to FUND regions.* Total
175 aggregate number of additional annual deaths (D_r) were calculated using age-specific population data
176 per country ($P_{i,a}$), and then summed across countries to yield results for the 9 FUND regions
177 presented in Tol (2002a) and the world (Online Resource 1 Table S7 for definition of regions):

$$178 \quad D_r = \sum_i \sum_a P_{i,a} M_{i,r,a} \quad (6)$$

179 Following Tol (2002a) we first assumed that heat only affects the urban population whereas cold
180 affects the entire population. In a second step, we relaxed this assumption and calculated additional
181 deaths from heat exposure based on total population data.

182 This step is important because recent epidemiological research suggests that rural populations might
183 even be more vulnerable to heat than urban populations (Shi et al. 2015). We are aware of the fact that
184 applying the same relationship to both urban and rural population introduces an additional error,
185 because vulnerabilities are known to be different. However, since our main goal has been to test the
186 sensitivity of FUND's parameterization to specific assumptions made in the underlying analyses by
187 Martens (1998) and Tol (2002a), rather than providing a recalibration of FUND, we think our
188 approach albeit simplistic is valid for the given purpose.

189 All data and programming code (Python 2.7.11) is available as Online Resource 2.

190 3 Results

191 Comparing the original temperature-mortality relationships of Martens (1998) with the results of our
192 reanalysis (Fig. 2) shows that the two flaws we identified in his data handling have mainly two
193 effects. First, Martens' decision to use data for all ages to parameterize age-group specific functions
194 introduce a bias towards lower relative mortality change per 1°C warming in the age group > 65
195 years, especially in the warm range (Fig. 2b). Second, Martens' questionable way of extracting
196 estimates from Green et al. (1994) leads to the overestimation of combined mortality change in the

197 cold range in both age groups (compared to our reanalysis based on modified estimates from Green et
198 al. 1994; Fig. 2).

199 In accordance, Martens would have projected considerable higher increases in heat-related mortality
200 in the elderly (> 65 years) in nearly all of the 17 countries considered (Fig. 3) if he had excluded all
201 ages data from the corresponding part of his meta-analysis. The results of our reanalysis also clearly
202 show, that a considerable part of the reductions in cold-related mortality in the age group > 65 years
203 reported by Martens (1998) are due to his misinterpretation of Green et al. (1994). The latter can also
204 be seen in age group < 65 years, but differences between Martens' (1998) original projections and our
205 reanalysis are generally smaller.

206 Next, we aimed at reproducing the extrapolation of Martens country-specific results to all countries
207 and the subsequent aggregation to FUND regions as presented by Tol (2002a). Since the regression
208 equations we derived based on Martens' original projections were not exactly the same as the
209 equations given in Tol (2002a) (Fig. 4, Online Resource 1 Table S6), we used both sets of equations
210 to extrapolate data to all countries. The corresponding aggregated mortality estimates at 1°C global
211 warming for the 9 FUND regions and the world are qualitatively very similar to the data reported by
212 Tol (2002a) (Fig. 5 a-c): Reduction in cold-related cardiovascular mortality overcompensate increases
213 in heat-related cardiovascular mortality in all regions, with a considerable net reductions in annual
214 deaths at the global level (in FUND approximately 380,000 fewer deaths annually, Online Resource 1
215 Table S8)

216 In order to investigate how the bias revealed in Martens (1998) translates into regional mortality
217 projections underlying FUND's damage functions, we calculated an additional set of regression
218 equations based on our reanalysis of Martens (1998), applying the two corrections simultaneously
219 (Fig. 4) (for supplementary results, applying the corrections one at a time, see Online Resource 1 Fig.
220 S4). The corresponding mortality projections for FUND region are quite different to the mortality data
221 currently used in the model: increases in heat-related cardiovascular mortality now outweigh
222 reductions in cold-related cardiovascular mortality in 4 out of 9 regions, with a much lower net
223 balance of approximately 90,000 fewer deaths per year (Fig. 5d, Online Resource 1 Table S8).

224 In a last step, we tested the sensitivity of Tol's regional numbers to his assumption on the complete
225 resistance of rural populations to heat-related cardiovascular stress. Using our reanalysis of Martens
226 (1998) and now also assuming that heat affects both urban and rural populations, we find stronger
227 absolute increases in heat-related mortality than reductions in cold-related mortality in all regions but
228 CEE&FSU and OECD-E (Fig. 5e). Globally the net balance has now shifted from negative to positive
229 with approximately 150,000 additional deaths annually (Online Resource 1 Table S8).

230 *4 Discussion and conclusions*

231 In this study, we set out to reproduce the empirical data basis on temperature-related cardiovascular
232 mortality, entering FUND and other economic assessment tools relying on the same calibration
233 scheme (e.g., ENVISAGE). Following Martens (1998) and Tol (2002a) to the extent possible, our
234 estimates matched FUND's regional mortality projections for 1°C global warming with respect to a
235 considerable net benefit of global warming due to strong reductions in cold-related mortality. Yet,
236 addressing two flaws identified in the data handling of Martens (1998) and one questionable
237 assumption made by Tol (2002a) indicated that both studies are likely biased towards finding a net
238 reduction of mortality under global warming – a bias that is transmitted into the damage functions of
239 FUND, ENVISAGE and related economic assessment tools.

240 Since we consider cardiovascular mortality only, while FUND also accounts for heat-related
241 respiratory diseases, one could argue that our results on net mortality effects cannot be directly
242 compared to FUND. Therefore, in a supplementary analysis we included the regional estimates on
243 heat-related respiratory mortality presented by Tol (2002a) in the calculation of net effects (Online
244 Resource 1 Fig. S4). As would be expected we find differences in absolute numbers and on the
245 regional level, but the major qualitative difference between FUND and our reanalysis (a net reduction
246 contrary to a net increase in global additional deaths) remains. It is also important to note that FUND
247 currently ignores cold-related respiratory diseases, despite their albeit inferior to cardiovascular
248 diseases but still non-negligible contribution to all-cause cold-related mortality (Ebi&Mills 2013).

249 One obvious conclusion from our study is that the damage functions on temperature-related mortality
250 to date employed in FUND and ENVISAGE need to be urgently replaced. Hsiang (2016) recently
251 outlined methodological requirements for empirical studies to measure the climate effect on economic
252 and social outcomes (including human health), potentially providing the basis for improved
253 calibrations of IAMs. In light of these requirements, both econometric and epidemiological research
254 on temperature-related mortality has made major advances since the publication of Martens (1998)
255 almost twenty years ago. In particular, improved methods now better account for temporal
256 displacement (i.e., ‘harvesting’, and delayed cold effects) inherent in temperature-mortality
257 relationships (Deschênes and Greenstone 2011, Gasparrini 2014). Due to the assemblage of large
258 datasets on temperature-related mortality in cities (counties) across geographical regions and the
259 development of customized meta-analytical approaches (Gasparrini et al. 2015a, Gasparrini et al.
260 2015b, Nordio et al. 2015) the first knowledge is emerging on factors that may explain across space
261 and time differences in observed exposure-response curves. For example, there is now evidence that
262 the increasing usage of air condition has strongly attenuated heat effects on mortality in the United
263 States over time (Nordio et al. 2015, Barreca et al. 2016).

264 It has also been firmly established that populations across countries and world regions are
265 acclimatized to their local climate, with MMTs dependent on latitudes or mean annual temperatures
266 (Guo et al. 2014, Tobias et al. 2017). Similarly, recent literature confirms the finding of earlier studies
267 (Kalkstein & Greene 1997) that due to acclimatisation current heat-mortality relationships are weaker
268 in warmer cities (Barreca et al. 2016), translating into relatively smaller heat impacts expected under
269 climate change (Schwartz et al. 2015). Furthermore, there is some indication that deviations of
270 temperatures in the cold range exert stronger effects in regions with relatively mild climates (Guo et
271 al. 2014, Huang et al. 2015, Nordio et al. 2015). Interestingly, this recent evidence on temperature-
272 related mortality across climate zones provides an additional argument for the urgent need of updating
273 FUND’s damage functions. In fact, the regression equations used by Tol (2002a) and in this study to
274 extrapolate mortality projections to countries worldwide imply that the cooler (warmer) the country
275 today the greater the expected reduction (increases) in cold(warm)-related mortality in the future.
276 Thus, what follows from the equations is exactly the opposite of what would be expected based on the
277 most recent literature (for the supplementary discussion of further conceptual caveats regarding
278 Martens (1998) and Tol (2002a) see Online Resource 1 section B).

279 So, given the important advances since the publication of Martens (1998), why is it not
280 straightforward to provide a new empirical data basis on temperature-related mortality for
281 recalibrating FUND and other global scale IAMs? Importantly, even in the most comprehensive
282 recent studies on temperature-related mortality across space (covering almost 400 cities in 13
283 countries on 4 continents, Gou et al. 2014, Gasparrini et al. 2015a) there are major data gaps
284 remaining (e.g., all of Africa). It is controversially discussed whether one should extrapolate known
285 exposure-response relationships to these unexplored locations differing strongly in terms of health

286 care systems and infrastructure (Deschênes 2014). It is also important to note that while studies such
287 as Gou et al. (2014) and Gasparrini et al. (2015a) represent a major step forward in our understanding
288 of *current* temperature-mortality relationships across space, this new data has so far not been used to
289 project temperature-related mortality under different scenarios of climate change. To our knowledge
290 there is at present only one globally-gridded dataset on temperature-related mortality projections
291 (Honda et al. 2014, WHO 2014), with the important limitations that one single exposure-response
292 curve derived from Japanese cities is employed across the global, that cold-related mortality is
293 neglected entirely, and that only the age group > 65 years is considered. Last but not least, in order to
294 derive damage functions to be integrated into IAMs, mortality projections will need to be analysed as
295 a function of different degrees of GMT rise (James et al. 2017), instead of the more conventional
296 approach of analysing future shifts in mortality along the dimension of time (see, e.g., recent studies
297 by Schwartz et al. 2015, Lee and Kim 2016, Guo et al. 2016). With the exception of one study on
298 temperature-related years of life lost in Brisbane, Australia (Huang et al. 2012), we are not aware of
299 any other study that has attempted to specifically derive the relationship between differential changes
300 in GMT and corresponding changes in temperature-related mortality.

301 In conclusion, we show that the mortality projections of Martens (1998) and their extension by Tol
302 (2002a) have found their way into several recent economic assessments of climate change, most
303 prominently through the application of the widely used IAMs FUND and ENVISAGE. Therefore, the
304 bias for which we provide evidence here is relevant also for the results of these assessments, such as
305 estimates of the SCC or economy-wide damage costs. Although temperature-related mortality is only
306 one of the many climate impact channels to be considered, the likely overestimation of benefits of
307 climate change in this sector would obviously translate into the underestimation of total climate
308 damage costs. Thus, it would be expected that the comparatively low global damage estimates derived
309 from FUND and ENVISAGE (Revesz et al. 2014) will be revised towards higher costs, once these
310 models rely upon improved global projections for cold- and heat-related mortality.

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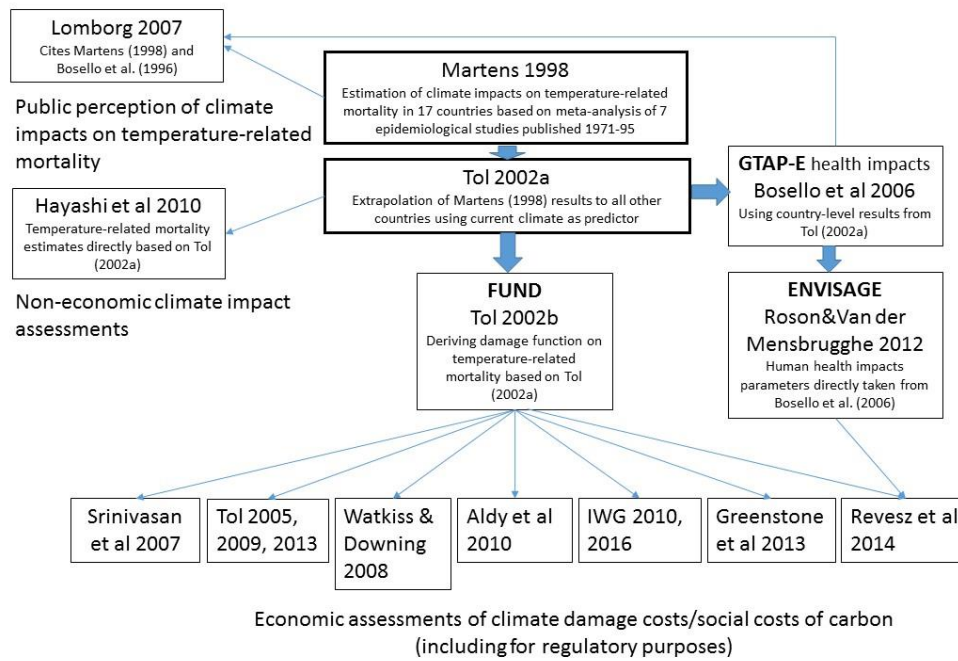
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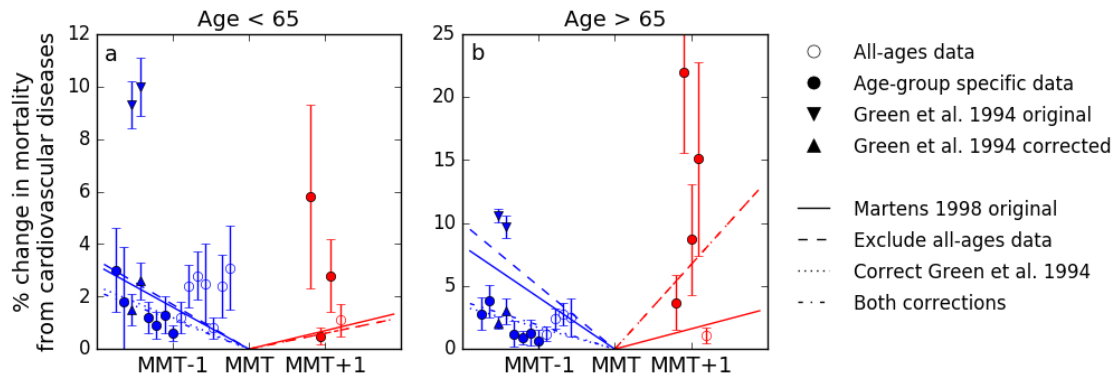
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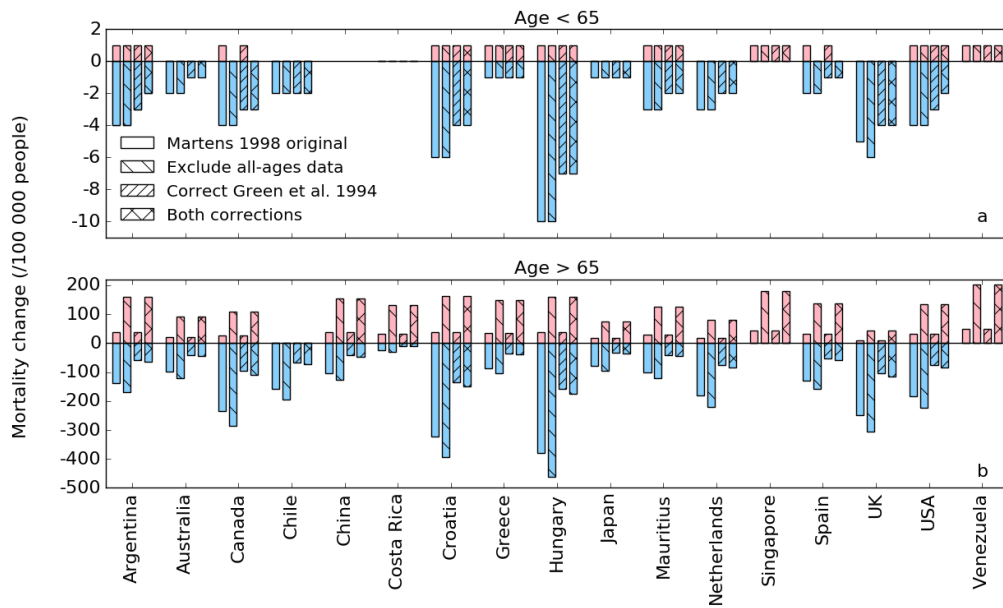
423 **Fig. 1** Use of temperature-related mortality projections by Martens (1998) and Tol (2002a) in
 424 economic (and non-economic) assessments of climate damages, including integrated assessment
 425 modelling with the models FUND and ENVISAGE.

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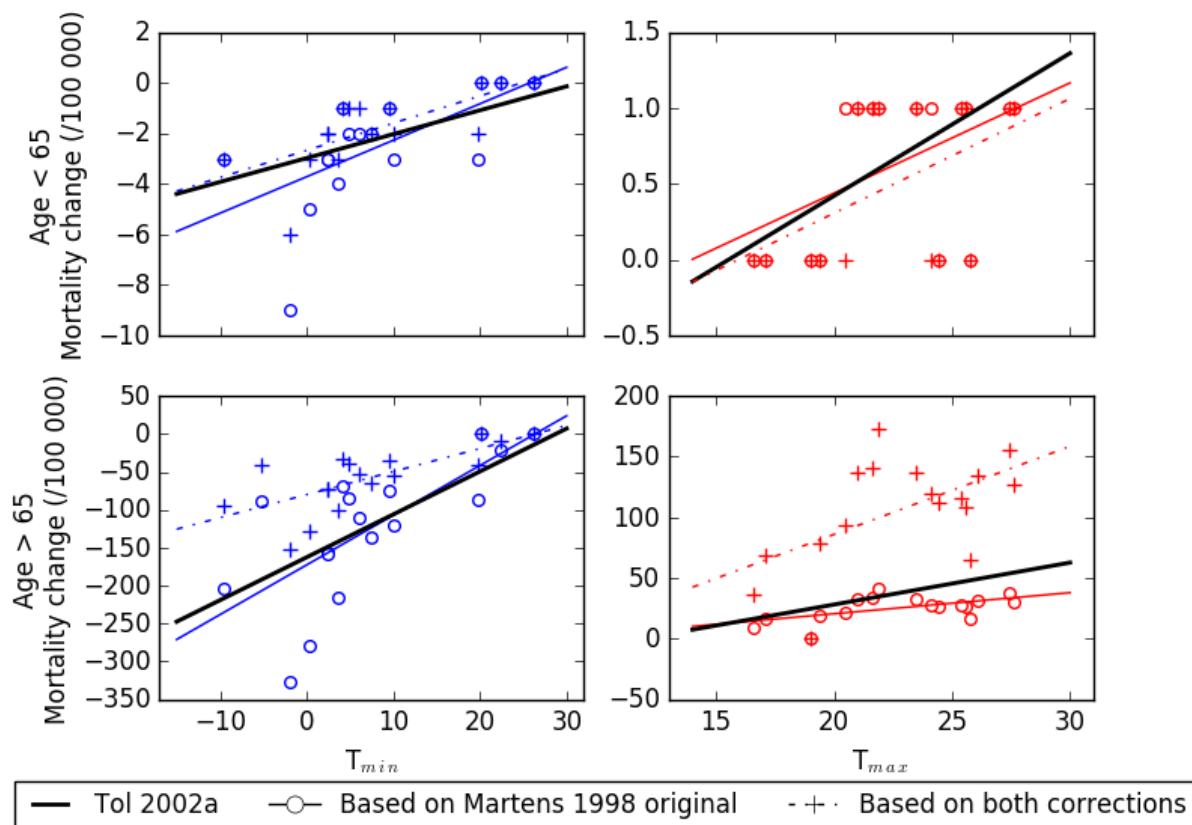


427

428 **Fig. 2** Temperature-mortality relationships from Martens (1998) (solid lines) and revised versions
 429 from our reanalysis (broken lines) for a) age group < 65 years, and b) > 65 years. Symbols show
 430 individual estimates $\beta_j (\pm s_j)$ of changes in cardiovascular mortality for an increase of 1°C in the cold
 431 range (blue, temperature < MMT) and in the warm range (red, temperature > MMT), used to calculate
 432 combined effect estimates $\bar{\beta}$ (and $\tilde{\beta}$), i.e., to parameterize the slopes (see Eq. 1). Note that data points
 433 are displaced along the x-axis for better visibility, and that slopes are *not* regression lines.

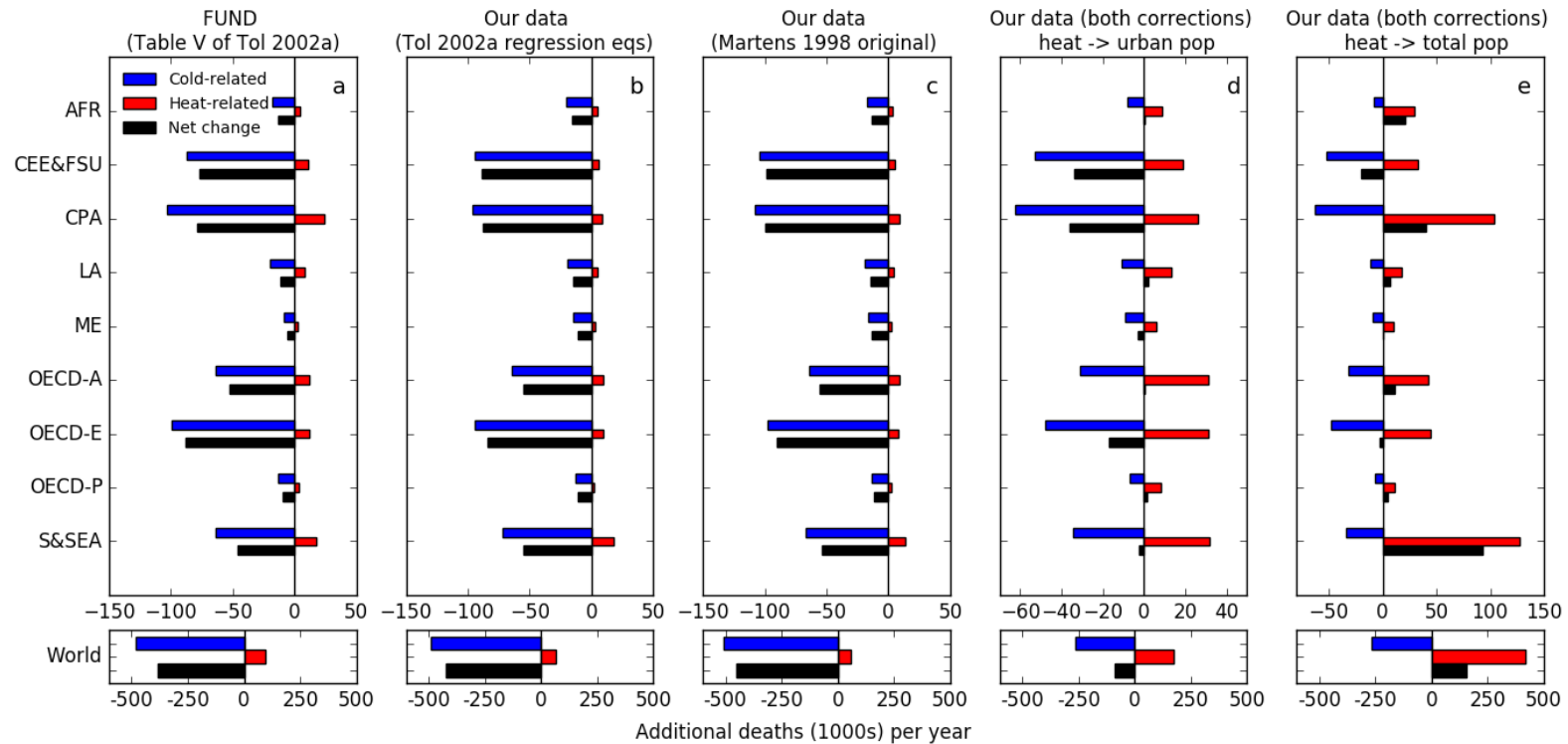


434
 435 **Fig. 3** Annual changes in people dying due to cardiovascular diseases at approximately 1.2°C global
 436 warming (a) < 65 years old and (b) > 65 years old in 17 countries based on Martens (1998). Decreased
 437 cold-related mortality (temperature < MMT, see Fig. 2) is shown in blue, increased heat-related
 438 mortality (temperature > MMT) in red. Plain bars depict original estimates of Martens (1998), hatched
 439 bars show results based on revised temperature-mortality functions (cf. broken lines in Fig. 2)



440

441 **Fig. 4** Linear regressions used by Tol (2002a) and in our reanalysis to extrapolate cold (blue) and heat
 442 (red) related changes in cardiovascular mortality for age groups < 65 and > 65 years (cf., Fig. 3) to
 443 countries worldwide, as a function of minimum and maximum monthly mean temperatures in capital
 444 cities (T_{min} , T_{max}). For equations see Online Resource 1 Table S6



445

446 **Fig 5** Regional (top) and global (bottom) number of additional annual cardiovascular deaths (1000s) at 1°C increase in global mean temperature (a) as
 447 reported by Tol (2002a) and (b-e) as estimated based on regression equations shown in Fig. 4 (see also Online Resource 1 Table S6). (a-d) is based on the
 448 assumption that heat affects only urban populations as in the FUND model; (e) is based on the assumption that both urban and rural populations are affected.
 449 Region abbreviations: AFR: Africa, CEE&FSU: Central and Eastern Europe and the former Soviet Union, CPA: Centrally planned Asia, LA: Latin America,
 450 ME: Middle East, OECD-A: OECD - America (excl. Mexico), OECD-E: OECD Western Europe, OECD-P: OECD Pacific (excl. South Korea), S&SEA:
 451 South and Southeast Asia (see Online Resource 1 Table S7 for list of countries)