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Mortality Impact of Low Annual Crop Yields in a Subsistence Farming Population of Burkina Faso under the Current and a 1.5°C Warmer

3

Climate in 2100

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- 14

15 Abstract

In subsistence farming populations of sub-Saharan Africa reliant on rainfed agriculture, years of low 16 17 crop yields result in poorer child nutrition and survival. Estimates of such impacts are critical for their 18 reduction and prevention. We developed a model to quantify such health impacts, and the degree to which they are attributable to weather variations, for a subsistence farming population in the Nouna 19 20 district of Burkina Faso (89,000 people in 2010). The method combines data from a new weather-crop 21 yield model with empirical epidemiological risk functions. We quantify the child mortality impacts 22 for 1984–2012 using observed weather data and estimate potential future burdens in 2050 and 2100 using daily weather data generated by global climate models parameterized to simulate global 23 24 warming of 1.5°C above pre-industrial levels. For 1984–2012, crop yields below 90% of the period 25 average were estimated to result in the total of 109.8 deaths per 10,000 children <5 years, or around 26 7,122.0 years of life lost, 72% of which are attributable to unfavourable weather conditions in the crop growing season. If all non-weather factors are assumed to remain unchanged, the mortality burden 27 28 related to low crop yields would increase about twofold under 1.5°C global warming by 2100. These 29 results emphasize the importance and value of developing strategies to protect against the effects of 30 low crop yields and specifically the adverse impact of unfavourable weather conditions in such 31 settings under the current and future climate.

32 Keywords: climate change, health, child mortality, crop yield, agriculture, vulnerable population

33 **1. Introduction**

Studies from Ethiopia, Mali, and Burkina Faso (Belesova et al., 2017a; Grace et al., 2016; Johnson and Brown, 2014; Yamano et al., 2005) suggest that low crop yields are an important risk factor for child nutrition and health in subsistence farming populations of sub-Saharan Africa, and that low crop yields in the year of birth have an adverse effect on child survival (Belesova et al., 2017b; Johnson and Brown, 2014). To date, there have been few attempts to quantify these impacts at population level.

Annual variations in crop yields may arise from range of factors, including the weather conditions
during the growing season, pests and disease, changes in crop selections and management, food prices
and trade. In this paper, we focus on weather variations (precipitation, temperature, solar radiation) as
the principal factor of importance for a subsistence farming population dependent on rain-fed
agriculture (Ray et al., 2015).

45 Altered weather patterns under climate change may present additional challenges for such populations (Phalkey et al., 2015). Nelson et al (2009) suggested that climate change-related impacts on yields 46 may increase the worldwide number of underweight children by 24% by 2050, while Lloyd et al 47 48 (2018, 2011) suggests a 45% increase in child stunting in West Africa by the same year through 49 changes in crop productivity and an even further increase through changes in food prices and incomes 50 of the poorest 20% of the population. However, these estimates are based on aggregate-scale modelling rather than direct empirical relationships between the weather, food production, nutrition, 51 52 and health. A few empirical studies have analysed the association of variations in precipitation or in the Normalized Difference Vegetation Index (NDVI) with child nutrition and survival (Grace et al., 53 54 2016; Johnson and Brown, 2014; Kumar et al., 2016; Rukumnuaykit, 2003; Yamano et al., 2005), but these studies do not enable a clear interpretation of specific pathways through which these 55 56 associations operate.

57 In this paper, we report a method of quantification that links a (local) weather-crop yield model with 58 empirical epidemiological evidence to model the impact of low crop yields on child mortality in a 59 subsistence farming population of sub-Saharan Africa. The Paris Agreement of the United Nations 60 Framework Convention on Climate Change set the aspirational target to limit the global warming to a 61 1.5°C increase in global temperatures above preindustrial levels (UNFCCC, 2015). Evidence on the 62 potential climate change impacts under a scenario of meeting this target is synthesised in the special report of the Intergovernmental Panel on Climate Change (IPCC, 2018). Although the report 63 64 suggested that health risks associated with food insecurity are lower under 1.5°C than 2°C global warming, the possible magnitude of such health risks under a 1.5°C warmer climate remains unclear, 65 66 particularly in vulnerable populations. Here, we examine the potential effect of changes in weather patterns on crop deficits and their associated child mortality in a vulnerable subsistence farming 67 population of Burkina Faso under climate change consistent with a 1.5°C increase in global 68 69 temperatures above preindustrial levels.

70 **2.** Methods

71 **2.1. Study population and setting**

We developed a model of child mortality using data and evidence for the population of the Nouna 72 Health and Demographic Surveillance System (HDSS) in Kossi province, North West Burkina Faso. 73 The area of residence of this population is classified as dry orchard savannah with annual average 74 75 precipitation of 796 mm over the past five decades (Diboulo et al., 2012). In 2010, the Nouna HDSS 76 system covered a population of 89,000 from 59 villages located across a third of the area of Kossi 77 province (Schoeps et al., 2014). This population has been followed up by the Centre de Recherche en 78 Santé de Nouna (CRSN) since 1992 through regular surveys of demographic, socio-economic, and 79 health data (Sankoh and Byass, 2012). The single agricultural production season in this area covers the rainy season months of June to October (Dabat et al., 2012). The population relies almost 80 exclusively on subsistence farming based on rain-fed agriculture (Dabat et al., 2012). Irrigation has 81 82 been implemented in only 210 ha (<0.03%) of the 732,800 ha of the total area of the province (Dabat et al., 2012). 83

84 **2.2. Modelling methods**

We developed a model with two main elements: (1) a model of weather-crop yield relationships, and(2) a model of crop yield-mortality (Figure 1):

87 (i) Weather-crop yield model

We developed a separate weather-crop yield model for each of the five main crops grown in Kossi
province following the methods of Gornott and Wechsung (2016) adjusted to the specific local
conditions.

The models were estimated with log-transformed functional form (Cobb–Douglas production function) [eq 1] (Lee et al., 2013; You et al., 2009). We applied fixed effects transformation to the endogenous variable crop yield (y) and the vector of exogenous variables (x): G weather variables (g = 1, ..., G), H economic variables (h = 1, ..., H), here only acreage), and I (range 0 - 2) dummy variables (i = 1, ..., I) – used to control for extreme yield anomalies, as in Albers et al. (2017) and Blanc (2012).

97 [eq 1]

98
$$\log \ddot{y}_{jt} = \sum_{g=1}^{G} \beta_{jg} \log \, \ddot{x}_{jgt} + \sum_{h=1}^{H} \beta_{jh} \log \, \ddot{x}_{jht} + \sum_{i=1}^{I} \beta_{ji} \, x_{jit} + \log \ddot{u}_{jt}$$

 $\log \ddot{y}_t = \log \left(\frac{y_t}{\bar{y}}\right)$, \bar{y} as arithmetic average of y_t and respectively for x and u. The term β represents 99 the parameters, u is the error term, t as the time-index (t = 1, ..., T) and j the crop (millet, sorghum, 100 101 maize, fonio, and rice). Crop-specific yield anomalies, i.e., the ratio of the annual yield to the period average $(\ddot{y} = \left(\frac{y_t}{\bar{y}}\right))$, 1984–2012, were regressed on growing season weather parameters (Mainardi, 102 2011; Rowhani et al., 2011). Crop-specific relative yield estimates and acreage (ha) for Kossi 103 province were obtained from the national Annual Agricultural Surveys of Burkina Faso, which used a 104 consistent year by year province-level estimation approach based on the crops cut method (Direction 105 Générale des Prévisions et des Statistiques Agricoles, 2013). The crops cut estimation method is 106 107 potentially subject to bias (e.g., resulting from uneven plant density of the fields) (Fermont and 108 Benson, 2011). We minimised the risk of such bias by using measures of relative crop yield variation 109 from one year to another.

110 The growing season weather variables used in each crop model were selected from knowledge of plant physiology (Gornott and Wechsung, 2016; Schauberger et al., 2017). The variables were derived 111 from WFDEI^{*} data, originally available at 0.5° (approximately 50 km at the equator) spatial and daily 112 temporal resolution (Weedon et al., 2014). We extracted the data for the area of Kossi province and 113 114 derived relevant annual weather variable values by aggregating the weather data over the growing season of each crop (see Appendix A for technical details on weather variable derivation). We 115 assumed the following crop growing seasons for the Kossi province, as informed by data from the 116 117 local agricultural authority, the agricultural crop calendar of the Food and Agriculture Organisation, 118 and the Global Yield Gap and Water Productivity Atlas: for maize and millet 1 August to 31 October, for rice 1 May to 31 October, for sorghum 15 March to 31 October, and for fonio 15 April to 15 119 September (Food and Agriculture Organization, 2010; Yield Gap, 2013). 120

121 The following growing season variables were tested and retained in the model if they contributed to the model goodness of fit (R²): solar radiation, cumulative precipitation, mean vapour pressure deficit, 122 123 growing degree days (optimum temperature for crop growth of 8–30°C), killing degree days (temperature >30°C), days without precipitation, dry spells longer than 5 days, and heavy 124 precipitation events (>40 mm per day). In addition, total acreage under cultivation was included to 125 capture inter-annual changes in agricultural management (Iizumi and Ramankutty, 2015).Further 126 technical details on weather-crop yield models, including key assumptions of these models and 127 validation details, are provided in Appendix A. 128

Using these models we determined for each year of the analysis period the weather-attributable variation in each of the five main crop types (Albers et al., 2017), and hence the contribution of weather impacts to the annual yield deficit. To extract the weather-attributable share of the yield variability we applied the equation [1], setting acreage to constant, and exponentiating the result:

133 [eq 2]

^{*}WFDEI -- <u>Water and Global Change Forcing Data Methodology Applied to the European Centre for Medium-Range Weather Forecasts Interim Re-Analysis</u>

134
$$y_{jt} = \exp\left(\left(\sum_{g=1}^{G} \beta_{jg} \log \ddot{x}_{jgt} + \sum_{h=1}^{H} \beta_{jh} \overline{\log \ddot{x}_{jht}} + \sum_{i=1}^{I} \beta_{ji} x_{jit}\right) + \overline{\log y_{jt}}\right)$$

135 *(ii) Crop yield-mortality model*

To estimate mortality impacts we used life tables (Miller and Hurley, 2003) constructed for the Nouna 136 137 population using age- and sex-specific Nouna HDSS mortality data, 1992-2012. A previous 138 epidemiological study provided relative risks for mortality in children under 5 years in relation to the 139 annual crop yield deficits for the year of the child's birth (Belesova et al., 2017b). The relative risks in 140 that study were derived as hazard ratios from Cox proportional hazard models adjusted for a range of 141 potential confounders. The confounders were determined a priori based on existing studies and the 142 local context, and included age, sex, season, crop yield in subsequent years of life, ethnicity, religion, 143 mother's and father's ability to read, semi-rural vs rural residence, indicators of village infrastructural 144 characteristics (presence of a market, health care facility, drilled water wells, and quality of road connection), other village-level random effects, time trend (calendar year), and the existence of an 145 146 undernutrition treatment programme. The relative risks for child mortality in relation to the annual crop yield deficits for the year of the child's birth is most likely a consequence of in utero exposures 147 148 to poor nutrition on the part of the mother leading up to birth, or of poor nutrition during the first year 149 of life (Belesova et al., 2017b). Risks associated with exposure to crop deficits at subsequent years of 150 life or exclusively in utero were not available from the existing literature for the purpose of our life 151 table models.

The crop yield deficit was defined by a parameter we refer to as the annual Food Crop Productivity Index (FCPI) (Belesova et al., 2017b). The FCPI reflects the weighted average of the yield (kg/ha) of the five main food crops in the Nouna area (millet, sorghum, maize, fonio (a form of millet), rice) relative to their annual mean yield for the period of 1992–2012, and is expressed as a percentage of the period average. Thus, an FCPI of 80% represents a 20% deficit in overall food crop yield in Kossi province relative to the period average. Algebraically, for year i, the FCPI was calculated as follows:

159

$$FCPI_i = \sum_{j=1}^{J} y_{ij} * w_{ij}$$

- 161 FCPI_i relative food crop yield (%) for year i
- 162 y_{ij} yield of crop j in year i relative to its mean yield in 1992–2012
- 163 w_{ij} harvest of crop j in year i as a proportion of the total harvest across the five food crops
- 164 j identifier of each food crop (millet, sorghum, maize, fonio, rice) with j = 1, ..., J

The needed input data to compute the FCPI – the annual crop harvests (kg), acreage (ha), and yields
(kg/ha) for each of the five food crops – were obtained from national Annual Agricultural Surveys
supplied by the Agricultural Statistics Service of Burkina Faso (Direction Générale des Prévisions et
des Statistiques Agricoles, 2013).

169 With these data, we were therefore able to derive the annual crop deficits and hence compute the

170 relative risk for mortality of children born in the same year up to 5 years of age. The relative risk of

171 mortality in a given year was applied to the cohort of children born in that year and to their mortality

172 risk in each subsequent year until the age of 5 years but not at older ages. Applying this relative risk to

- the life table gave a calculation of the change in number of child deaths and years of life lost (YLL).
- 174 Such calculations were done for years with a crop yield of 90% or lower than the period average.

175 Years with yields greater than 90% of the period average were assumed to carry no excess risk of

176 mortality.

The FCPI was also converted into the weight of grain per adult equivalent per year (kg/ae/year) and its food energy value (Stadlmayr et al., 2012) per adult equivalent per day (kcal/ae/day). We made this conversion using evidence from another study (Belesova et al., 2017a), where those values were available.

181 **2.3.** Quantifying the impact of crop yield deficits

Using the modelling methods outlined above, we carried out two sets of computations of the childmortality impacts associated with:

184 (1) The *observed* pattern of crop deficits for the period 1984–2012 and the fraction
185 attributable to unfavourable growing season weather conditions during this period; and

186 (2) A 'thought experiment' of *projected* future crop deficits for 2050 and 2100 under a
187 pattern of climate change consistent with a 1.5°C warming above pre-industrial levels
188 (but assuming all other factors are held constant at baseline levels).

To examine how crop yields might vary in the *future* under climate change, we applied the regression 189 coefficients from the crop models (calibrated on the observed weather and yield data) to weather 190 191 projections data derived from two general circulation model (GCM) realizations provided by the 192 Inter-Sectoral Impact Model Inter-comparison Project (ISI-MIP2b), an international climate-impact 193 modelling network, for the period of 1700-2100: IPSL-CM5A-LR (Institute Pierre Simon Laplace Climate Model) and MIROC5 (Model for Interdisciplinary Research on Climate) (Frieler et al., 2016), 194 195 corrected against reanalysed weather observations by the ISI-MIP. These two model realizations were chosen because they cover a wide range of the projected uncertainty in future changes in precipitation 196 197 in our study region. Both model realizations project warmer future temperatures, while MIROC5 198 projects wetter conditions and IPSL-CM5A-LR projects dryer conditions. Our estimates do not 199 account for future changes in any other factors than weather (e.g., agricultural management practices 200 and adaptation, prices, socio-economic and demographic conditions etc.) to indicate the independent 201 effect of climate change alone. The projections were summarized for 30 year periods centred on 2015 202 (current conditions), 2050, and 2100. The GCM realizations used in this paper correspond to a 203 conservative assumption of a global mean temperature increase of 1.5°C above pre-industrial levels 204 by the end of the century (Frieler et al., 2016), the aspirational target agreed at the 2015 Paris 205 conference of the UNFCCC (UNFCCC, 2015).

Additional analyses were performed based on upper and lower bounds of the model parameters as away of exploring the influence of parameter uncertainty on the results (Appendix B).

8

208 The study was conducted following the ethical standards of the Declaration of Helsinki and was

209 approved by the [name of the institution blinded for peer-review] Observational Ethics Committee and

210 the Comité Institutionnel d'Ethique du Centre de Recherche en Santé de Nouna.

3. Results

212 **3.1.** Crop yield deficits and mortality impacts, 1984-2012

Over this 29 year period, crop yields were <90% of the period average in eight years, and <80% in two years (Figure 2, Table 1). The yield deficit in years with <90% of the period average yield was equivalent to an annual average harvest deficit of 18.2 kg/ae/year (kilograms/adult equivalent/year) or 177.9 kcal/ae/day – when averaged across all 29 years of observation. The lowest annual crop yield, 65% of the period average, was observed in the year 2000. It was equivalent to a harvest deficit of 109.8 kg/ae/year or 1,073.5 kcal/ae/day in that year, which is more than a third of the recommended daily food energy intake (2,900 kcal/ae/day) for a moderately active adult male of 30–60 years of age.

The impact attributed to these crop yield deficits was estimated as 3.8 deaths or 245.6 YLL per 10,000
children under 5 years, when averaged across all 29 years of the period of observation (1984–2012) –
Table 1. For the year of lowest crop yield (2000), the attributed mortality impact was equivalent to
22.9 deaths or 1,477.3 YLL per 10,000 children under 5 years.

Figure 3 shows the cumulative total of the attributed child mortality and YLL over our period of observation 1984–2012, reaching the total of 109.8 child deaths and 7,122.0 YLL per 10,000 children under 5 years. Over this period, *weather* factors during the crop-growing season were estimated to account for 72% of the crop deficit and mortality impacts (Table 1). Over the period of observation, the additional mortality from crop yields <90% of the period average represented around 1.45% (95% CI 0.29, 8.61%) of all-cause mortality in children <5 years in that period.

230 **3.2.** *Projected crop yields deficits and mortality under 1.5* °C global warming

231 Figure 4 shows the crop deficit simulations from the GCM models with future projections of weather 232 data. Crop yield projections demonstrate some background variability (which is the likely reason for the difference in the direction of FCPI projections between the two GCMs around 2050 in Table 2). 233 Despite the background variability, the outputs of both GCMs suggest progressively less favourable 234 growing conditions by the middle and end of the 21^{st} century, with a steeper decline projected by 235 IPSL-CM5A-LR than MIROC5. This is mainly because of an increase in temperatures above the 236 optimal levels for crop growth and changing precipitation variability (Figure 5). As expected, both 237 GCMs suggest an increase in the number of days with temperatures above the optimal levels for crop 238 growth, and decreasing precipitation by IPSL-CM5A-LR but increasing precipitation by MIROC5 by 239 the end of the century. 240

241 The per cent of years with FCPI <90% of the baseline period average, and the corresponding mortality 242 impacts, were estimated approximately to double from the period of 2000-2030 to the period of 243 2085–2115 (Table 2). Central estimates suggested that the annual average mortality impact attributed to crop yield deficits could raise between 2000–2030 and 2085–2115 from 3.8 to 5.8 child deaths and 244 245 from 245.7 to 374.3 YLL per 10,000 children under 5 years, according to the IPSL-CM5A-LR climate model, or from 1.6 to 3.3 child deaths and from 102.5 to 210.0 YLL per 10,000 children under 5 246 247 years, according to the MIROC5 climate model. Further analyses suggested that there is considerable 248 uncertainty in these estimates. Results based on the lower bound of the parameters were compatible 249 with no effect in any of the projection time periods, while results based on the upper bound of the 250 parameters suggested a possibility that the annual average child mortality impact of crop yield deficits might reach 99.3 child deaths and 6,388.1 YLL per 10,000 children under 5 years in the year 2100, 251 252 assuming all other factors than weather remain constant (see Appendix B: Supplementary figures and 253 tables).

254 **4. Discussion**

The evidence we present here provides, to our knowledge, the first empirically grounded estimates of the impact of low crop yields, and weather-related low crop yields, on child mortality in a subsistence farming population of sub-Saharan Africa.

The estimates reflect the evidence of an adverse effect of low crop yield in the year of birth on child survival to the age of five years. They suggest an appreciable impact of current crop yield variations and patterns of changing climate that are likely to increase those impacts even under the very conservative assumptions of 1.5°C global warming. Although based on models and data specific to the Nouna HDSS population of Kossi province, Burkina Faso, our findings are likely to be broadly indicative of the impact of low crop yields in other similar populations in the region.

264 Our model of health impact was based only on the mortality impacts of low yields, as the only outcome with a clearly established relationship with annual crop yield deficits. Furthermore, yield 265 deficit in the year of birth was the only timing of exposure examined in relation to a health outcome, 266 267 as relative risks for other exposure timings were not available in the literature. The relative risks used 268 in this study were estimated by relating child survival from birth (to 5 years of age) to relative yield of 269 the last harvest preceding or at the time of the date of birth (Belesova et al., 2017b). Hence, these 270 relative risk estimates reflect the effect of exposures to crop yield deficits that children in the study 271 population experienced in part in utero and in part in their first year of life. The modelling inputs 272 available for our analyses do not allow estimating the child survival effect of exposures to crop yield 273 deficit beyond the first year of life or of exposures that occurred exclusively in utero. Our results may 274 not capture the full health burden of the cumulative lifetime exposure to low crop yields. Other 275 epidemiological evidence suggests that in utero exposure and low crop yields in later childhood, 276 adolescence, and adulthood may also have negative effects on health and survival (Belesova et al., 277 2017b). Furthermore, our estimates did not consider morbidity impacts related to undernutrition, the associated increased susceptibility to infectious diseases, compromised cognitive development and 278 279 immunity, as well as compromised productivity in later life (Belesova et al., 2017b). Currently, the epidemiological evidence is too insecure to allow the development of a health impact model that 280 integrates all these effects, but they may add appreciably to our current estimates. Furthermore, we 281

defined the deficits in relation to the period average crop yield level, which might be sub-optimal for the nutritional needs of this population (Belesova et al., 2017a). Therefore, we interpret our estimates as representing a lower bound estimate of the actual impact of crop yield deficits on child health.

285 Our estimates of future impacts of low crop yields under climate change should not of course be 286 interpreted as showing the real impact to be expected in the future. Many factors other than climate are likely to change over the course of this century (and some have already changed in the past, e.g., 287 288 child mortality rate, crop yield productivity, and population size) which will have a bearing on the 289 health and survival of the child population in subsistence farming areas of sub-Saharan Africa. 290 Exploring whether changes in such factors have made a contribution in the past and are likely to contribute in the future to increased or decreased vulnerability of our study population was beyond the 291 292 scope of our study. The purpose of our estimates was rather to indicate the likely influence of changes 293 in weather patterns alone if all other factors are held constant. With this assumption, they suggest that 294 climate change would have an appreciably adverse impact even using the conservative rise of 1.5°C 295 above preindustrial levels.

296 Our future projections of weather parameters, crop yields, and child mortality differed depending on 297 the GCM. Our study area is located in a region where GCM agreement on changes in precipitation is 298 relatively low (Schewe and Levermann, 2017). MIROC5, which projects wetter conditions, suggested 299 somewhat smaller reduction in the projected future crop yields and correspondingly smaller increase 300 in the attributed child mortality by 2100, as compared to IPSL-CM5A-LR, which projects drier future conditions (when compared to the current levels). In climate impact assessments, it is important to 301 302 cover a range of climate model outputs. The two GCMs used in our assessment reflect relatively high 303 and low estimates in changes in temperature and precipitation (Schewe and Levermann, 2017). 304 Despite the background variability in crop yield projections, our results suggest consistent trends of decline in future crop yield projections across both GCMs. The central estimates of our projections 305 under both GCMs suggested an increase in the attributed child mortality by 2100 with the 1.5°C 306 307 global temperature increase above preindustrial levels, as compared to the current conditions. Future

research should attempt further disentangling the variability and uncertainty in yield and childmortality projections using a wider range of GCMs.

Current nationally determined contributions of all signatories of the Paris agreement would lead to a 3°C warming, unless radical emission reduction is undertaken by all countries (IPCC, 2018). The impacts are likely to be greater under other climate change scenarios projecting a greater extent of global temperature increase than 1.5°C (Blanc, 2012) and requires further investigation. In this sense, our projections of the yield deficit attributed child mortality also give a lower bound estimate of the impact of climate change on child mortality in subsistence farming populations. Our results support the urgent need to limit any further global warming to 1.5°C above the pre-industrial levels.

317 It was beyond the scope of this paper to consider what form such strategies or interventions might take. Determining this is a complex scientific undertaking and requires assessment of a range of 318 factors and implementation research. However, as our weather-crop yield model shows, low crop 319 320 yields are largely attributable to the variations in the weather (Blanc, 2012). Moreover, applying the 321 model to daily weather data generated under the 1.5°C global warming target suggests that crop yields will fall over time (all other things being equal) as growing season temperatures rise and the 322 323 distribution of precipitation becomes less reliable. This observation suggests that efforts targeted 324 specifically at ameliorating the effects of weather on crop yield should be considered (e.g., use of 325 drought resistant seeds or improved irrigation as risk reduction measures or crop insurances as risk 326 transfer measures). Nutritional protection measures and interventions such as food and supplement distribution, conditional cash transfers, food-for-work programmes, crop insurance schemes or other 327 support might also be appropriate (del Ninno et al., 2005). 328

As with any modelling study, there are uncertainties and limitations. Several of our modelling inputs, including the central exposure–response (yield–mortality) function, which was derived from a single relevant study available to date (Belesova et al., 2017b), and modelled crop yield based on assumptions detailed in Appendix A, were based on limited data leading to limitations concerning the precision of the function and modelled yield estimates. Further uncertainty concerns the precipitation 334 trends projected by the GCMs, as explained above. Other general circulation models might have yielded somewhat different estimates of changes in weather patterns, and we deliberately did not 335 attempt to account for the (uncertain) trends in non-climate factors, such as socio-economic 336 development. Other studies reporting climate impact projections at the national, and global levels 337 338 attempted accounting for such trends using the country-level projections of Gross Domestic Progress (GDP), population, education, and urbanisation available under the Shared Socioeconomic Pathways 339 (SSPs) scenario framework (Riahi et al., 2017). Socio-economic and demographic scenarios are not 340 yet available at sub-regional level of Burkina Faso and hence were beyond the scope of our analyses. 341 Nonetheless our analyses offer valuable first insights into the magnitude of the child mortality 342 burdens related to crop yield deficits. It is noteworthy that the full health and wellbeing burden of the 343 cumulative lifetime exposure to low crop yields is likely to be appreciably greater than our estimated 344 345 impact.

346 Future research on health impacts of crop yield variation in the context of weather variability should attempt to address a broader set of health outcomes and wider aspects of their temporal effects, 347 beyond the effect of the exposure in the year of birth. To strengthen the evidence and provide 348 conclusive policy advice, similar modelling studies are required on other settings with high 349 350 prevalence of rain-fed subsistence agriculture. Such studies require further epidemiological evidence 351 on the exposure-response function of crop yield variation and health outcomes from other settings, as well as meta-analytical estimates of this function across settings once more empirical studies are 352 available. Furthermore, improvements in the precision of the exposure-response function and crop 353 354 yield modelling, which require long time-series yield data, as well as advancements in the reduction 355 of uncertainty associated with the general circulation models are important steps for future research.

5. Conclusion

This study contributes evidence of an appreciable impact of low crop yields on population health in the subsistence farming population of rural Burkina Faso. Much of this health impact appears to be related to the negative agricultural impact of increasing weather variability, which is likely to worsen under climate change (all other factors being equal). These results emphasize the importance and
value of developing strategies to protect against the effects of low crop yields and specifically the
adverse impact of unfavourable weather conditions during the growing season on crop yields in such
settings.

Tables

	Ave	rage year*	Worst year (2000)		
	Overall	Weather-related	Overall	Weather-related	
Deficit in food crop harvest and yield					
kg per adult equivalent/year	18.2	13.5	109.8	106.2	
kcal per adult equivalent/day	177.9	131.8	1,073.5	1,038.4	
% FCPI below the period average	6%	4%	35%	34%	
Mortality per 10,000 children <5 years					
Child deaths (<5 years)	3.8	2.9	22.9	21.9	
YLL (not discounted)	245.6	185.7	1,477.3	1,418.3	

 Table 1. Crop deficits and weather-related crop deficits, and their attributed child mortality impact, 1984–2012.

*Based on deficits in years with FCPI<90% averaged across all years of the period 1984–2012.

Table 2. Crop yield, and attributable mortality impacts under 1.5°C global warming.

	IPSL-CM5A-LR			MIROC5		
	2015 (2000–2030)	2050 (2035–2065)	2100 (2085–2115)	2015 (2000–2030)	2050 (2035–2065)	2100 (2085–2115)
No (%) years with yield (FCPI) relative to the period average of:						
<90%*	11 (37%)	7 (23%)	17 (57%)	5 (17%)	13 (43%)	11 (37%)
<80%	1 (3%)	3 (10%)	2 (7%)	0 (0%)	0 (0%)	1 (3%)
Average annual deficit in food crop harvest						
kg per adult equivalent/year	18.3	13.6	28.0	7.6	18.6	15.8
kcal per adult equivalent/day	179.1	133.4	273.9	74.8	181.9	154.4
Deaths per 10,000 children <5 years						
Average year**	3.8	2.9	5.8	1.6	3.8	3.3
Worst year	16.9	15.3	16.2	12.1	12.2	15.5
YLL per 10,000 children <5 years						
Average year**	245.7	185.0	374.3	102.5	246.9	210.0
Worst year	1,086.9	980.0	1,038.9	774.1	785.8	994.1

*Includes years with FCPI<80%. **Based on deficits in years with FCPI<90% averaged across all years in the respective 30-year time periods.

Figures



Figure 1. Flowchart of weather-agriculture-health modelling approach.

To develop estimates of child mortality attributable to low crop yields under the current and a 1.5°C warmer climate we combined agricultural models of weather effect on crop yields (components: observed crop yield and acreage data, observed and projected weather data, crop calendar) with the calculations of child mortality impact attributable to annual crop yield deficits (components: age- and sex-specific mortality rates, published risk ratio of child mortality impact of low crop yields (Belesova et al., 2017b), observed and projected crop deficits).

Abbreviations: FCPI, Food Crop Productivity Index; WCM, Weather-Crop Model, WA-FCPI, Weather-Attributed Food Crop Productivity Index; YLL, Years of Life Lost.



Figure 2. Time series of the Food Crop Productivity Index (FCPI) with the corresponding mortality risk ratio.

Circled markers indicate years with FCPI<90%. Arrows with dotted lines show the extent of the yield deficit below the counterfactual of 100% FCPI.



Figure 3. Cumulative health impact incurred over the period of 1984–2012 and attributed to the exposure in the year of birth to the overall and weather-attributed yield deficit in years with FCPI<90%.



Figure 4. Food Crop Productivity Index (FCPI) projections based on climate data of each of the general circulation models separately: A – IPSL-CM5A-LR, B – MIROC5. Red colouring at the bottom of the time series indicates years when FCPI declines below 90%.

Linear trends: (1) black – trend in FCPI, suggesting -0.45 (95% CI -0.90, 0.01) and -0.20 (95% CI -0.63, 0.24) percentage point change in FCPI per decade for IPSL-CM5A-LR and MIROC5, respectively.



Figure 5. Future projections of selected the key weather variables explaining crop yield variability (killing degree days and precipitation over the growing season of millet – the main food crop in Nouna area) and their correlation with the Food Crop Productivity Index (FCPI), presented for each Global Circulation Model.

Linear trends: increase in killing degree days by 0.03 (95% CI 0.01, 0.06) °C per decade and a decrease in precipitation of -4.40 (95% CI -9.34, 0.54) mm, according to IPSL-CM5A-LR; increase in killing degree days by 0.04 (95% CI -0.08, 0.15) °C per decade and a decrease in precipitation of -58.81.40 (95% CI -289.59, 171.98) mm, according to MIROC5.

Appendix A: Technical details on weather-crop yield models

1. Weather variable derivation

Variable	Unit	Purpose	Calculation
solar radiation	J/cm ²	to determine crop	sum over the growing season
(SR)		growth potential	
precipitation	mm	to capture deviations	sum over the growing season
(PREC)		from the optimal	
Vapour pressure	mm	to capture the	sum of daily vapour pressure deficit values over the
deficit (VPD)	111111	atmospheric water	growing season, as derived from the maximum
		demand	(TMP_{max}) and minimum temperature (TMP_{min})
			(DVWK, 1996, Sonntag, 1990, Roberts et al., 2013)
			$UDD = \left(11 \left(2 \left(\frac{17.269 TMP_{\text{max}}}{237.3 + TMP_{\text{max}}} \right) 2 \left(\frac{17.269 TMP_{\text{min}}}{237.3 + TMP_{\text{max}}} \right) \right)$
			$VPD = 0.11 \left(e^{(257.5 + 1)M1 \max^2} - e^{(257.5 + 1)M1 \min^2} \right)$
growing degree	°C	to explain the	sum of days with daily mean temperature falling within
days (GDD)		(positive) influence	the range of optimal temperature for the growing
		on crop growth	season, 30–8 °C for all examined crops
killing degree	°C	to account for	cumulated temperature sum of daily mean temperature
days (KDD)		temperatures leading	above the optimal temperature (of 30 °C) over the
		potentially negative	growing season
		impact on crop yields	
		(Roberts et al., 2013)	
days without	days	to capture	sum of days with no precipitation over the growing
precipitation		precipitation	season, identified as follows:
(DWP)		distribution which	$DWD = \sum_{dum}^{D} dum = (1, \text{ if } PREC_d = 0)$
		crop development	$DWP_t = \sum_{d=1}^{d} awp_d = \{0, \text{ if } PREC_d > 0\}$
		erop de terophient	d = the day within each of the crop development
			periods $(d = 1, \dots, D)$
dry spells longer	days	to capture crop yield	number of days with dry spells longer than 5 days over
than 5 days (SP5)		impact of the dry	the growing season, identified as:
		spens	
			$= \sum SP5_d = \begin{cases} 1 & \text{if } RD_d \ge 5 & \& RD_{d+1} = 0\\ 0 & \text{if } RD_d \ge 5 & \& RD_{d+1} \neq 0 \end{cases}$
			$(RD_{d-1} + 1)$ if $PREC_d < 0.5$)
			with $RD_d = \begin{cases} 1.5 & a^{-1} \\ 0 & \text{if } PREC_d \ge 0.5 \end{cases}$ and RD
1			as rainy day
precipitation	of the	impact of soil erosion	number of events over the growing season, identified
events >40mm	events	and nitrogen leaching	$\frac{D}{D}$ (4.16 DDE G > 00.)
per day (PE40)			$PE_t = \sum PE_d = \begin{cases} 1 & \text{if } PREC_d \ge 20 \\ 0 & \text{if } PREC_d \le 20 \end{cases}$
			$\underbrace{\begin{array}{c} & & \\ & &$
acreage	ha	to capture changes in	hectares of land cultivated under the respective crop
		agronomic	type in Kossi province
		practices and land use	
		(lizumi and	
		Ramankutty, 2015) in	
		the model	

Table A1. Weather variables used for the construction of the crop-specific statistical models.

2. Crop yield model assumptions

The crop models were based on the following assumptions:

(1) We assumed the relationship of weather and management impacts on crop yields to be linear Since we use a statistical regression model with only linear exogenous variables, non-linear yield impacts were not considered, which corresponds to the approach of Schlenker & Lobell (2010). To ensure that our models did not omit such impacts, we conducted a statistical test (RESET) (Croissant and Millo, 2008).

(2) We assumed that weather variables have equal impact on yield at every stage of crop development The magnitude of the effect of weather variation on crop yield, in terms of grain quantity and quality, differs depending on the stage of crop development during which the crop was exposed to these weather variations (Rötter and Van de Geijn, 1999). However, many statistical crop models, e.g., the models of Moore & Lobell (2014), Blanc (2012) and You et al (2009) did not divide the growing period into sub-periods to allow for differential impact of weather variables in these sub-periods and showed that weather variables aggregated over the entire growing season are able to sufficiently explain crop yield variability. We used the out of sample cross validation to corroborate robustness of our crop models (in which weather variables were aggregated over the entire growing season) for yield estimation beyond the time period of the observed yield data. The out of sample cross validation confirmed the robustness of our models.

(3) We assumed that estimated model parameters are valid for the future climate conditions of 1.5 °C warming

Estes et al (2013) and Lobell & Burke (2010) show that statistical models have high capacity to reproduce observed conditions (often better than process-based models), however, they are more limited in their ability to project in unobserved conditions. As the future climate conditions under 1.5 °C global warming may be relatively similar to the current climate conditions, we assumed that our model parameters are valid for these conditions. A comparison of the past and future climate data showed that the range of the inter-annual weather variability observed in the past included most of the variability projected under the 1.5 °C of global warming.

(4) We assumed that fixed effects transformation controls for any time-invariant effects, such as soil conditions, market access, and land tenure, which we assumed to be time-invariant

Our model captures time-invariant effects like the soil conditions or other farm-specific conditions though the fixed effects variable transformation (Wooldridge, 2013). The fixed effect transformation eliminates time-invariant effects in the data by capturing them implicitly in the statistical model. We assumed that the transformation allows controlling for such factors as investment in agricultural equipment, market access, land tenure security, and soil conditions (Brasselle et al., 2002; Lay et al., 2009). Under these assumptions, we suggest that the model parameters are not biased by these time-invariant effects (no omitted variable bias).

(5) Management impacts are reflected by crop acreage

Often, information on agronomic management is not available for many regions in sub-Saharan Africa (Müller and Robertson, 2014). Since changes in acreage are often an indicator for changes in soil quality and available labour, we used acreage to capture possible effects of such factors (Iizumi and Ramankutty, 2015; Wouterse, 2010).

3. Additional considerations in crop model validation

(1) The model goodness of fit for each crop type is shown in table A2:.

 \mathbb{R}^2 Variables Crop Fonio GDD, KDD, DWP, Acreage_Fonio, dummy84, dummy85 0.92 KDD, SR, VPD, SP05, Acreage_Maize Maize 0.51 Millet 0.64 PREC, GDD, KDD, VPD, SR, SP05, PE40, Acreage_Millet, dummy00, dummy90 GDD, KDD, VPD, SR, DWP, SP05, PE40, Acreage_Rice Rice 0.53 PREC, GDD, KDD, VPD, SR, DWP, Acreage_Sorghum, dummy00, dummy90 Sorghum 0.54

Table A2. Model goodness of fit and variables for each crop type.

(2) The weather variables explained large parts of crop yield variability. In comparison to the full model (in parenthesis), the weather variables explained the following percentage of yield variability: 50% (51%) for maize, 86% (92%) for fonio, 63% (64%) for millet, 51% (53%) for rice, and 54% (55%) for sorghum. Although the effect of the acreage is rather small, it was retained in the model as

a measure of reducing the risk of bias. In Kossi, acreage of the respective crops shows strong interannual changes and long-term increase in fonio and millet by 64% (10-year averages) and 36% respectively. The maize and sorghum acreage declined by 57% and 45%. Rice acreage shows a very strong increase of +337%, but this is mostly driven by few observations in the mid-1990s and from 2008 to 2010 much above the average level. Yet, mostly the variable acreage shows no significant contribution to the explained yield variability. Despite this strong inter-annual and long-term change, we concluded that land productivity was unlikely to have changed in this period and that the farmers have not moved to less suitable land.

(3) We conduct several statistical tests to verify model robustness and validity. The statistical tests are described by Croissant & Millo (2008). The regression equation specification error test (RESET) was used to investigate whether quadratic variables are missed in the model. The RESET showed that quadratic variables were not neglected for any of the crops. The Breusch–Godfrey and Breusch–Pagan tests were applied to test against autocorrelation and heteroscedasticity. In two cases the model residuals were autocorrelated (fonio and millet), the other crops show no autocorrelation (Breusch–Godfrey test). As the time series are relatively short (T = 28) and the variable transformation tends to cause autocorrelation (Baltagi, 2005), we judged that this was unlikely to bias parameters in the models of fonio and millet. There appeared to be no heteroscedasticity (Breusch–Pagan test) in any of the models. The distribution of residuals was tested using the Shapiro–Wilk test, suggesting normal distribution of residuals in all models.

Appendix B: Supplementary figures and tables

Table B1 provides uncertainty estimates of the results reported in the Table 1 of the main text.

Table B1. Central and uncertainty estimates of the mortality impact of crop deficits and weather-related crop deficits over the period of 1984–2012.

Across columns and rows of the table, lower and upper bound estimates were based on different sources of uncertainty.

First and third columns (from the left): uncertainty estimates derived using the 95% confidence interval bounds (instead of the central estimate) of the risk ratio of child survival in relation to crop yield.

Second and fourth columns (from the left): uncertainty estimates derived using the 95% confidence interval bounds (instead of the central estimates) of the risk ratio of child survival in relation to crop yield and of the estimates of the weather-attributed part of crop yield variation.

	Averag	e year*	Worst year (2000)		
	Overall	Weather- related Overall		Weather- related	
Mortality per 10,000 children <5 years					
Child deaths (<5 years)	3.8	2.9	22.9	21.9	
	(0.7, 24.8)	(0.2, 25.5)	(4.6, 126.4)	(1.3, 126.1)	
YLL (not discounted)	245.6	185.7	1,477.3	1,418.3	
	(47.6, 1,591.7)	(5.5, 1,636.5)	(296.7, 8,169.6)	(84.2, 8,151.8)	

*Deficits in years with FCPI<90% averaged across all years of the period 1984–2012.

Table B2 provides uncertainty estimates of the results reported in the Table 2 of the main text.

Table B2. Central and uncertainty estimates of the annual average attributable years of life lost per 10,000 children <5 years under 1.5 $^{\circ}$ C global warming for three 30-year periods centred on years 2015, 2050, and 2100.

The uncertainty estimate are based on the lower and upper bounds of the 95% confidence intervals of the model parameters, i.e., mortality risk ratio and crop yield projections.

		IPSL-CM5A	-LR	MIROC5			
	2015	2050	2100	2015	2050	2100	
	(2000–2030)	(2035-2065)	(2085-2115)	(2000-2030)	(2035-2065)	(2085-2115)	
Deaths per 10 000 children <5 years							
Average vear*	3.8	2.9	5.8	1.6	3.8	3.3	
Average year	(0, 91.2)	(0, 97.0)	(0, 99.3)	(0, 82.4)	(0, 87.1)	(0, 91.2)	
Worst year	16.9	15.3	16.2	12.1	12.2	15.5	
worst year	(0, 156.8)	(0, 152.4)	(0, 138.9)	(0, 129.4)	(0, 131.7)	(0, 145.9)	
YLL per 10,000 children <5 years							
Average year*	245.7	185.0	374.3	102.5	246.9	210.0	
	(0, 5,869.2)	(0, 6, 234.5)	(0, 6, 388.1)	(0, 5, 300.8)	(0, 5, 600.0)	(0, 5,862.4)	
Worst year	1,086.9	980.0	1,038.9	774.1	785.8	994.1	
-	(0.10.068.1)	(0, 9.787.7)	(0, 8.919.4)	(0, 8.313.3)	(0, 8.462.0)	(0, 9.369.0)	

*Based on deficits in years with FCPI<90% averaged across all years in the respective 30-year time periods.

Note: the lower estimates in the projections are equal to 0 as a result of our assumption that mortality impact is incurred only in years with yield <90% of the baseline period average. The lower bound of the modelled crop yield estimates in all cases exceeded 90% FCPI, hence, not incurring mortality impact as a result of our modelling assumption that mortality impact is only incurred in years with FCPI <90%.



Figure B1. Cumulative mortality impact attributable to crop yield deficits in years with FCPI<90% (A) and to weather-related crop deficits (B).

Dashed lines represent uncertainty estimates based on 95% confidence intervals of the relative risk for child mortality used in the calculations (Belesova et al., 2017b).



Figure B2. Estimates of weather-attributable variation in crop yields, Kossi province, Burkina Faso, 1984–2012, based on estimates of the weather–crop model. Shaded bands indicate 95% confidence intervals.



Figure B3. Crop-specific yield and Food Crop Productivity Index (FCPI) projections based on climate data of each of the general circulation models separately: A – IPSL-CM5A-LR, B – MIROC5.

References

- Albers, H., Gornott, C., Hüttel, S., 2017. How do inputs and weather drive wheat yield volatility? The example of Germany. Food Policy 70, 50–61.
- Baltagi, B.H., 2005. Econometric Analysis of Panel Data, Third edit. ed. John Wiley & Sons Ltd.
- Belesova, K., Gasparrini, A., Sié, A., Sauerborn, R., Wilkinson, P., 2017a. Household cereal crop harvest and children's nutritional status in rural Burkina Faso. Environ. Heal. 16, 1–11.
- Belesova, K., Gasparrini, A., Sié, A., Sauerborn, R., Wilkinson, P., 2017b. Annual Crop Yield Variation, Child Survival and Nutrition among Subsistence Farmers in Burkina Faso. Am. J. Epidemiol. kwx241, 15–17.
- Blanc, E., 2012. The Impact of Climate Change on Crop Yields in Sub-Saharan Africa. Am. J. Clim. Chang. 01, 1–13.
- Brasselle, A.S., Gaspart, F., Platteau, J.P., 2002. Land tenure security and investment incentives: puzzling evidence from Burkina Faso. J. Dev. Econ. 67, 373–418.
- Castellvi, F., Perez, P.J., Stockle, C.O., Ibaiiez, M., 1997. Methods for estimating vapor pressure deficit at a regional scale depending on data availability. Agric. For. Meteorol. 87, 243–252.
- Castellví, F., Perez, P.J., Villar, J.M., Rosell, J.I., 1996. Analysis of methods for estimating vapor pressure deficits and relative humidity. Agric. For. Meteorol. 82, 29–45.
- Croissant, Y., Millo, G., 2008. Panel data econometrics in R: The plm package. J. Stat. Softw. 27.
- Dabat, M., Zongo, I., Kiendrebeogo, R., 2012. Etude sur les relations entre marchés et sécurité alimentaire au Burkina Faso. Montpellier: CIRAD.
- del Ninno, C., Dorosh, P.A., Subbarao, K., 2005. Food Aid and Food Security in the Short- and Long Run: Country Experience from Asia and sub-Saharan Africa. The World Bank Social Safety Nets Primer. [online] Available at: http://www.fao.org/fsnforum/sites/default/files/files/Have %20your%20say/Food_aid_in_Afri 427 ca_and_Asia.pdf [Accessed 14 Oct 2017]
- Diboulo, E., Sié, A., Röcklov, J., Niamba, L., Ye, M., Bagagnan, C., Sauerborn, R., 2012. Weather and mortality: retrospective analysis of 10 years of data from Nouna Health and Demographic Surveillance System, Burkina Faso. Glob. Health Action 5, 1–8.
- Direction Générale des Prévisions et des Statistiques Agricoles et Alimentaires (DPSAA), 2011. Resultats Definitifs de l'Enquete Permanente Agricole 2007/2008. Ouagadougou: DPSAA
- Dvwk (1996) Ermittlung der Verdunstung von Land- und Wasserflächen, Bonn, Wirtschafts- und

Verlagsgesellschaft Gas und Wasser mbH.Estes, L.D., Beukes, H., Bradley, B.A., Debats, S.R., Oppenheimer, M., Ruane, A.C., Schulze, R., Tadross, M., 2013. Projected climate impacts to South African maize and wheat production in 2055: a comparison of empirical and mechanistic modeling approaches. Glob. Chang. Biol. 19, 3762–3774.

- Fermont A, Benson T. Estimating yield of food crops grown by smallholder farmers: A Review in the Uganda Context. IFPRI Discussion Paper No. 01097. 2011;1–57.
- Food and Agriculture Organization, 2010. Crop calendar. [online] Available at: http://www.fao.org/agriculture/seed/ cropcalendar/welcome.do [Accessed 29 March 2017]
- Frieler, K., Betts, R., Burke, E., Ciais, P., Denvil, S., Deryng, D., Ebi, K., Eddy, T., Emanuel, K.,
 Elliott, J., Galbraith, E., Gosling, S.N., Halladay, K., Hattermann, F., Hickler, T., Hinkel, J.,
 Huber, V., Jones, C., Krysanova, V., Lange, S., Lotze-Campen, H., Mengel, M., Mouratiadou, I.,
 Müller Schmied, H., Ostberg, S., Piontek, F., Popp, A., Reyer, C.P.O., Schewe, J., Stevanovic,
 M., Suzuki, T., Thonicke, K., Tian, H., Tittensor, D.P., Vautard, R., van Vliet, M., Warszawski,
 L., Zhao, F., 2016. Assessing the impacts of 1.5°C global warming simulation protocol of the
 Inter-Sectoral Impact Model Intercomparison Project (ISIMIP2b). Geosci. Model Dev. Discuss.
 1–59.
- Gornott, C., Wechsung, F., 2016. Statistical regression models for assessing climate impacts on crop yields: A validation study for winter wheat and silage maize in Germany. Agric. For. Meteorol. 217, 89–100.
- Grace, K., Nagle, N.N., Husak, G., 2016. Can Small-Scale Agricultural Production Improve Children's Health? Examining Stunting Vulnerability among Very Young Children in Mali, West Africa. Ann. Am. Assoc. Geogr. 4452, 1–16.
- Iizumi, T., Ramankutty, N., 2015. How do weather and climate influence cropping area and intensity? Glob. Food Sec. 4, 46–50.
- IPCC, 2018: Summary for Policymakers. In: Global warming of 1.5°C. An IPCC Special Report on the impacts of global warming of 1.5°C above pre-industrial levels and related global greenhouse gas emission pathways, in the context of strengthening the global response to the threat of climate change, sustainable development, and efforts to eradicate poverty [V. Masson-Delmotte, P. Zhai, H. O. Pörtner, D. Roberts, J. Skea, P. R. Shukla, A. Pirani, W. Moufouma-Okia, C. Péan, R. Pidcock, S. Connors, J. B. R. Matthews, Y. Chen, X. Zhou, M. I. Gomis, E. Lonnoy, T. Maycock, M. Tignor, T. Waterfield (eds.)]. World Meteorological Organization: Geneva.

Johnson, K., Brown, M.E., 2014. Environmental risk factors and child nutritional status and survival

in a context of climate variability and change. Appl. Geogr. 54, 209-221.

- Kumar, S., Molitor, R., Vollmer, S., 2016. Children of Drought: Rainfall Shocks and Early Child Health in Rural India. Popul. Dev. Rev. 42, 53–68.
- Lay, J., Narloch, U., Mahmoud, T.O., 2009. Shocks, Structural Change, and the Patterns of Income Diversification in Burkina Faso. African Dev. Rev. 21, 36–58.
- Lee, B.H., Kenkel, P., Brorsen, B.W., 2013. Pre-harvest forecasting of county wheat yield and wheat quality using weather information. Agric. For. Meteorol. 168, 26–35.
- Lloyd, S.J., Bangalore, M., Chalabi, Z., Kovats, R.S., Hallegatte, S., Rozenberg, J., Valin, H., Havlík,
 P., 2018. A Global-Level Model of the Potential Impacts of Climate Change on Child Stunting
 via Income and Food Price in 2030. Environ. Health Perspect. 126, 1–15.
- Lloyd, S.J., Kovats, S.R., Chalabi, Z., 2011. Climate change, crop yields, and undernutrition: Development of a model to quantify the impact of climate scenarios on child undernutrition. Environ. Health Perspect. 119, 1817–1823.
- Lobell, D.B., Burke, M.B., 2010. On the use of statistical models to predict crop yield responses to climate change. Agric. For. Meteorol. 150, 1443–1452.
- Mainardi, S., 2011. Cropland use, yields, and droughts: Spatial data modeling for Burkina Faso and Niger. Agric. Econ. 42, 17–33.
- Miller, B.G., Hurley, J.F., 2003. Life table methods for quantitative impact assessments in chronic mortality. J. Epidemiol. Community Health 57, 200–6.
- Moore, F.C., Lobell, D.B., 2014. Adaptation potential of European agriculture in response to climate change. Nat. Clim. Chang. 4, 610–614.
- Müller, C., Robertson, R.D., 2014. Projecting future crop productivity for global economic modeling. Agric. Econ. 45, 37–50.
- Nelson, G.C., Rosegrant, M.W., Koo, J., Robertson, R., Sulser, T., Zhu, T., Ringler, C., Msangi, S., Palazzo, A., Batka, M., Magalhaes, M., Valmonte-Santos, R., Ewing, M., Lee, D., 2009. Climate Change and Agriculture Impacts and Costs of Adaptation, Food Policy Report. Washington, D.C.
- Phalkey, R.K., Aranda-Jan, C., Marx, S., Höfle, B., Sauerborn, R., 2015. Systematic review of current efforts to quantify the impacts of climate change on undernutrition. Proc. Natl. Acad. Sci. 112, E4522–E4529.
- Ray, D.K., Gerber, J.S., MacDonald, G.K., West, P.C., 2015. Climate variation explains a third of

global crop yield variability. Nat. Commun. 6, 1–9.

- Riahi, K., van Vuuren, D.P., Kriegler, E., Edmonds, J., O'Neill, B.C., Fujimori, S., Bauer, N., Calvin, K., Dellink, R., Fricko, O., Lutz, W., Popp, A., Cuaresma, J.C., KC, S., Leimbach, M., Jiang, L., Kram, T., Rao, S., Emmerling, J., Ebi, K., Hasegawa, T., Havlik, P., Humpenöder, F., Da Silva, L.A., Smith, S., Stehfest, E., Bosetti, V., Eom, J., Gernaat, D., Masui, T., Rogelj, J., Strefler, J., Drouet, L., Krey, V., Luderer, G., Harmsen, M., Takahashi, K., Baumstark, L., Doelman, J.C., Kainuma, M., Klimont, Z., Marangoni, G., Lotze-Campen, H., Obersteiner, M., Tabeau, A., Tavoni, M., 2017. The Shared Socioeconomic Pathways and their energy, land use, and greenhouse gas emissions implications: An overview. Glob. Environ. Chang. 42, 153–168.
- Roberts, M.J., Schlenker, W., Eyer, J., 2013. Agronomic Weather Measures in Econometric Models of Crop Yield with Implications for Climate Change. Am. J. Agric. Econ. 95, 236–243.
- Rötter, R., Van de Geijn, S.C., 1999. Climate Change Effects on Plant Growth, Crop Yield and Livestock. Clim. Change 43, 651–681.
- Rowhani, P., Lobell, D.B., Linderman, M., Ramankutty, N., 2011. Climate variability and crop production in Tanzania. Agric. For. Meteorol. 151, 449–460.
- Rukumnuaykit, P., 2003. Crises and Child Health Outcomes: The Impacts of Economic and Drought /Smoke Crises on Infant Mortality and Birthweight in Indonesia. Michigan State Univ.
- Sankoh, O., Byass, P., 2012. The INDEPTH network: Filling vital gaps in global epidemiology. Int. J. Epidemiol. 41, 579–588.
- Schauberger, B., Gornott, C., Wechsung, F., 2017. Global evaluation of a semi-empirical model for yield anomalies and application to within-season yield forecasting. Glob. Chang. Biol., 23, 4750–4764.
- Schewe, J., Levermann, A. 2017. Non-linear intensification of Sahel rainfall as a possible dynamic response to future warming, Earth Syst. Dynam., 8, 495-505.
- Schlenker, W., Lobell, D.B., 2010. Robust negative impacts of climate change on African agriculture. Environ. Res. Lett. 5, 1–8.
- Schlenker, W., Roberts, M.J., 2009. Nonlinear temperature effects indicate severe damages to U.S. crop yields under climate change. Proc. Natl. Acad. Sci. U.S.A. 106, 15594–8.
- Schoeps, A., Souares, A., Niamba, L., Diboulo, E., Kynast-Wolf, G., Müller, O., Sié, A., Becher, H., 2014. Childhood mortality and its association with household wealth in rural and semi-urban Burkina Faso. Trans. R. Soc. Trop. Med. Hyg. 108, 639–647.

- Sonntag D (1990) Important new Values of the Physical Constants of 1986, Vapour Pressure Formulations based on ITS-90, and Psychrometer Formulae. Meteorologische Zeitschrift, 4, 340-344.
- Stadlmayr, B., Charrondiere, U.R., Enujiugha, V.N., Bayili, R.G., Fagbohoun, E.G., Samb, B., Addy, P., Barikmo, I., Ouattara, F., Oshaug, A., Akinyele, I., George Amponsah Annor, K.B., Ene-Obong, H., Smith, I.F., Thiam, I., Burlingame, B., 2012. West African food composition table. FAO: Rome.
- UNFCCC, 2015. Adoption of the Paris Agreement: Proposal by the President Draft decision-/CP.21. UNFCCC: Paris.
- Weedon, G.P., Balsamo, G., Bellouin, N., Gomes, S., Best, M.J., Viterbo, P., 2014. Data methodology applied to ERA-Interim reanalysis data. Water Resour. Res. 50, 7505–7514.
- Wooldridge, J.M., 2013. Introductory Econometrics A modern approach, 5th ed. South-Western Cengage Learning: Mason.
- Wouterse, F., 2010. Migration and technical efficiency in cereal production: evidence from Burkina Faso. Agric. Econ. 41, 385–395.
- Yamano, T., Alderman, H., Christiaensen, L., 2005. Child growth, shocks, and food aid in rural Ethiopia. Am. J. Agric. Econ. 87, 273–288.
- Yield gap, 2013. Information from collected literature for Burkina Faso to be used for crop modelling for Burkina Faso. [online] Available at: http://www.yieldgap.org/gygamaps/pdf/Details%20on %20crop%20information%20to%20calibrate%20crop%20models%20for%20Burkina%20Faso. pdf [Accessed 14 October 2017]
- You, L., Rosegrant, M.W., Wood, S., Sun, D., 2009. Impact of growing season temperature on wheat productivity in China. Agric. For. Meteorol. 149, 1009–1014.